Hybrid Latent Class Measurement of Health States Lacking a Gold Standard

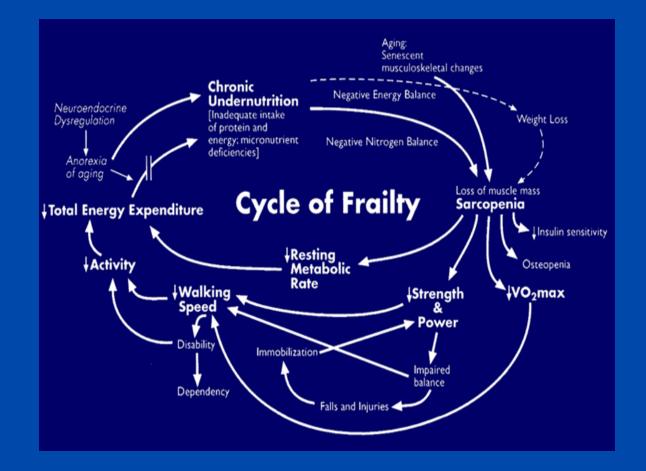
Karen Bandeen-Roche¹ Jeannie-Marie Sheppard² Jing Ning¹ Departments of Biostatistics¹ and Mental Health² Johns Hopkins Bloomberg School of Public Health

> ENAR Spring Meeting Atlanta, Georgia March 13, 2007

Measurement problem

- A health state
 - recognizable
 - conceptually well defined
 - has known consequences
- No gold standard
 - more than diagnostic error
 - no single consensus measurement
 - multifaceted consequences

Measurement Problem Geriatric Frailty



Fried et al., J Gerontol 2001; Bandeen-Roche et al., J Gerontol, 2006

Measurement Problem Aging

 Recognition - Chronic disease, disability, events - Variability among individuals • Theory: a biological process More than consequence accumulation – Multisystem dysregulation No gold standard - Even to the point of "surrogates"

Successful measures Classical Approach: Validity

Face: recognition

• Content: facets covered

• Criterion: utility

• Construct: theory – internal; external *DeVellis, 1991; Bartholomew, 1996*

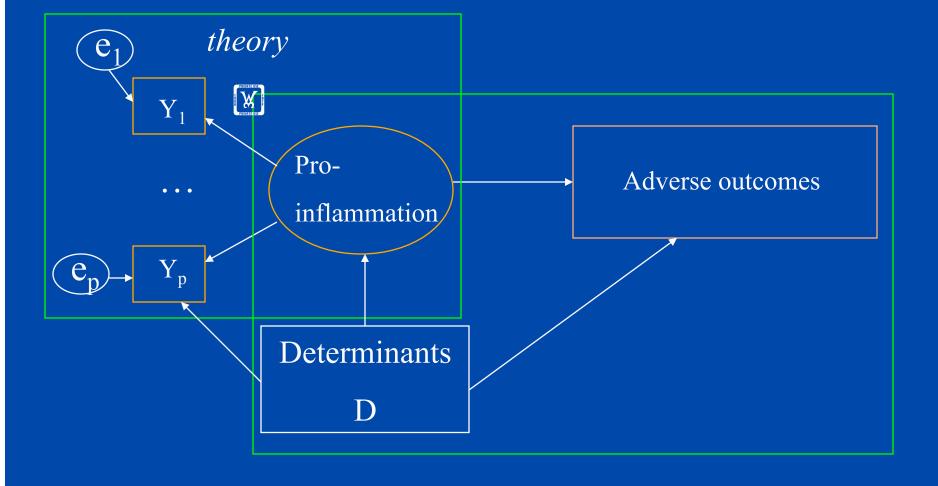
Points of the Introduction

- A well defined target; a less-welldefined operationalization
- Will retain validation for measure definition; performance evaluation
- Objective: Method to unify multiple validation aspects in 1 analysis

Outline

- Latent variable paradigm for measurement
- A new idea
 - Aims to balancing potentially conflicting validation premises
 - Application
- Discussion

Measurement Latent Variable Paradigm



Model

Kenneho

Generic

Specific (Latent Class Reg.; Categorical U=j , {1,...,J})

Measurement assumptions : $[Y_i|U_i, x_i]$

- conditional independence, nondifferential measurement

> heterogeneity in criterion presentation unrelated to measured or unmeasured characteristics

> fundamentally identifying

Latent Class Measurement How to obtain "indices"?

 Via posterior probabilities of class membership =

$$\hat{F}_{U|Y,x}(u \mid y, x)$$

• Then: exactly how?

- "Modal": by highest probability
- "Pseudo-classes": Randomize (Bandeen-Roche et al., 1997; Wang et al., 2005)

In what sense is LCA a "measurement" model?

Does it "discover" structure?

It operationalizes theory

 Science: Test if predictions borne out
 Most frequent theory: Homogeneity

Latent Class Measurement Syndrome Validation Application

Criteria manifestation is syndromic

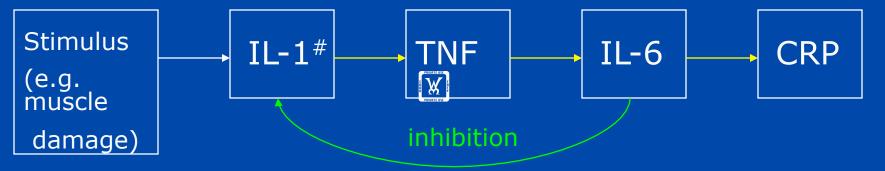
"a group of signs and symptoms that occur together and characterize a particular abnormality" (Webster Medical Dictionary 2003)

- If criteria characterize syndrome:
 - At least two clinically homogeneous groups (if <2, no co-occurrence)
 - No subgrouping of symptoms (otherwise, more than one abnormality characterized)
 Bandeen-Roche et al., J. Gerontol Med Sci, 2006

Measurement Application: Pro-Inflammation

- Central role: cellular repair
- A hypothesis: dysregulation key in adverse aging
 - Muscle wasting (*Ferrucci et al., JAGS 50:1947-54; Cappola et al, J Clin Endocrinol Metab 88:2019-25*)
 - Receptor inhibition: erythropoetin production / anemia (*Ershler, JAGS 51:S18-21*)

up-regulation



Difficult to measure. IL-1RA = proxy

Rationale of the New Work

- Which deserves pre-eminence?
 Internally validating assumptions?
 - Externally validating assumptions?
 Frailty: close tie to systemic dysregulation
 Depression: genetic "subtypes"
 Aging: tie to chronological age
 - Some compromise?

Rationale of the New Work

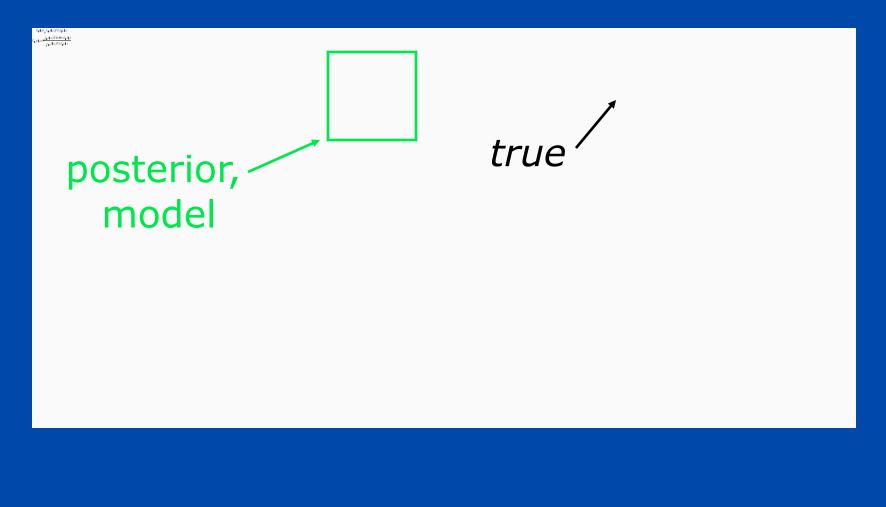
Which deserves pre-eminence?

 Internally validating assumptions
 Externally validating assumptions?
 Some compromise?

 A model (LCR) including externally validating variables and fitting by ML already "is" a compromise

A representation theorem

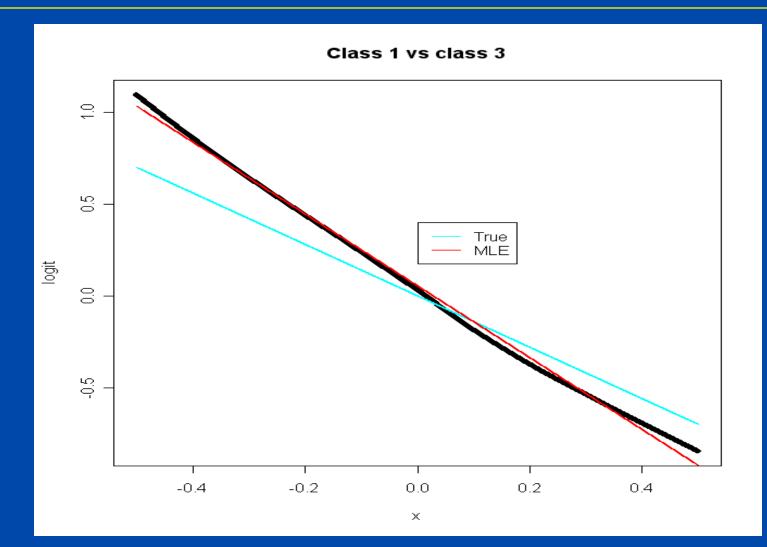
Consider "mixing" & "kernel" distributions:



A representation theorem

- Y_i is equivalent in distribution to Y* constructed as
 - 1) Generate V_i* from $F_{V|x}^{*}(v|x_i)$
 - 2) Given V_i*, generate Y* from $F_{Y|V,x}^*(y|V_i^*,x_i)$
- <u>Relevance</u>:
 - True for θ^* = Huber (1967) limit of MLE (e.g.)

True vs. realized mixing models



Rationale of the New Work

Which deserves pre-eminence?

 Internally validating assumptions
 Externally validating assumptions?
 Some compromise?

 Proposal: Allow stronger (or weaker) compromise than ML via "penalized" fitting

Implementing penalization

- <u>On LCR kernel</u>: Houseman, Coull & Betensky, *BMCS* online early
- <u>On LCR mixing distribution</u>: Sheppard Ph.D. thesis
- Key questions
 - Form of the penalty
 - Different purpose than usual?
 - What is the objective function?

Penalization Very brief background

• Fitting: minimize

-2 ln L(θ ;Y,x) + λ g(θ)

• Examples - "Ridge": $g(\theta) = \sum_{j} \theta_{j}^{2}$ - "Lasso": $g(\theta) = \sum_{j} |\theta_{j}|$

Green, Int Stat Rev, 1987; Tibshirani, JRSS-B, 1996

Penalization Very brief background

<u>A useful equivalence</u>: penalized fit obtains via formulating parameters as crossed random effects

 "Ridge": θ_j ~ N(0,σ{λ}²)
 "Lasso": θ_j ~ double exp(0,h{λ})

Wahba, JRSS-B, 1978; Ngo & Wand, J Stat Software, 2004

Form of the penalty Current case

Usual purpose: regularization

Here: secondary validation

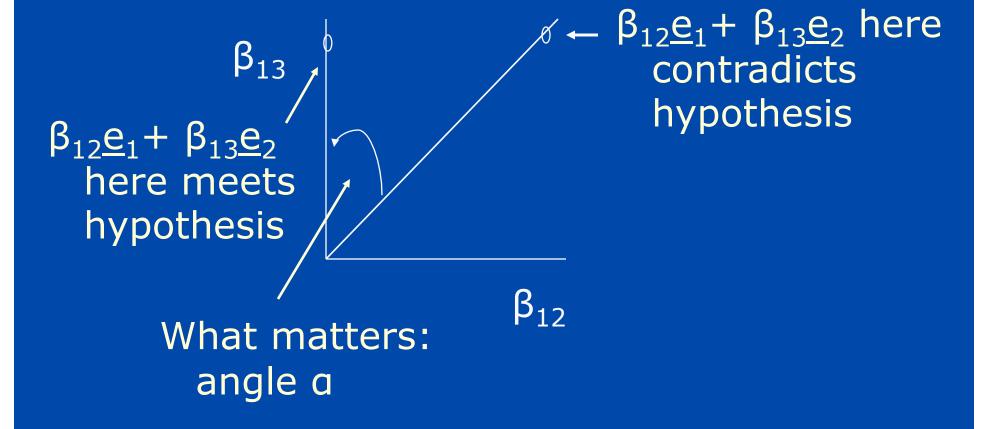
 <u>Discriminant hypothesis</u>: Genotypes predispose individuals to only one "subtype" of depression

- Say, LCR with one normal class (1) and two disordered classes (2, 3):
- Hypothesis: β_{1j} negligible, and $\beta_{1j'}$ appreciable, in

$$\log\left(\frac{p_k}{p_1}\right) = \beta_{0k} + \beta_{1k}x$$

with $p_k = pr(class k)$; x=genotype indicator; k=2,3; j, j' $\in \{2,3\}$; j \neq j'

Ridge, lasso not quite right



- Approach 1

 - Desired orientations are cos(a)=1, sin(a)=1
 - i.e., goal: minimize cos(a)+sin(a)

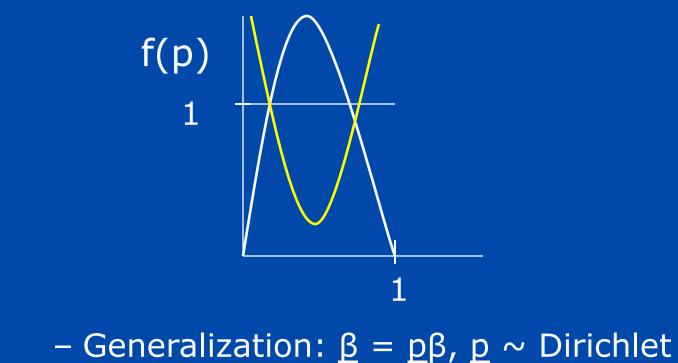
- i.e. minimize

$$\frac{|\beta_{12}| + |\beta_{13}|}{\sqrt{\beta_{12}^2 + \beta_{13}^2}}$$

• Approach 2

- Write $\beta_{12} = p\beta$; $\beta_{13} = (1-p)\beta$

Fit with beta random effect on p



Fitting Approach 2

• E-M algorithm: quite straightforward

- E-step: Computes posterior class membership probabilities given current parameter iterates
- M-step: minimize (e.g. Nelder-Mead)

 $-\sum_{i=1}^{n}\sum_{j=1}^{J}h(j|data)\ln[f_{U|x}(u|x,p,\beta)] + (1-\frac{\Delta}{2})\ln[p(1-p)]$

Simulation study Three-class model

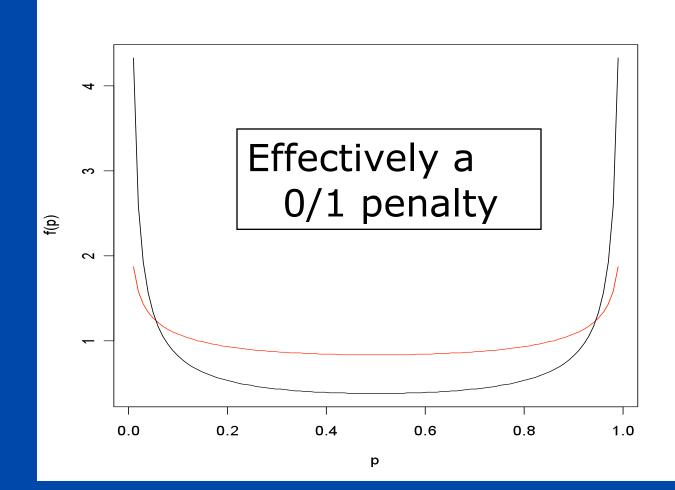
- 100 reps; single x~Unif(-.5,.5); n=1000
- Poly Log Reg: $\beta_{01} = \beta_{02} = 0$; $\beta_{13} = -1.4$; $\beta_{12} \in \{0, -0.5, -1.4\}$
- Measurement:

Class 1	Class 2	Class 3
.15	.85	.85
.15	.85	.85
.15	.85	.85
.15	.13	.85
.15	.13	.85

Simulation study Three-class model

- Two assumption scenarios
 Frank LCR
 - <u>Differential measurement</u>: First three items have increased log(odds =1) per unit x of 1.4 within each class

Simulation study Beta model: $\Delta = 1.5, .5$



Simulation Study Diff. Meas.; $\beta_{12}=0$; $\beta_{13}=-1.4$

Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β ₁₂	-0.04	0.14	-0.54	0.31
β ₁₃	-0.79	0.30	-1.01	0.34

Simulation Study Non-diff meas; $\beta_{12}=0$; $\beta_{13}=-1.4$

Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β ₁₂	0	0	0.04	0.32
β ₁₃	-1.42	0.35	-1.41	0.38

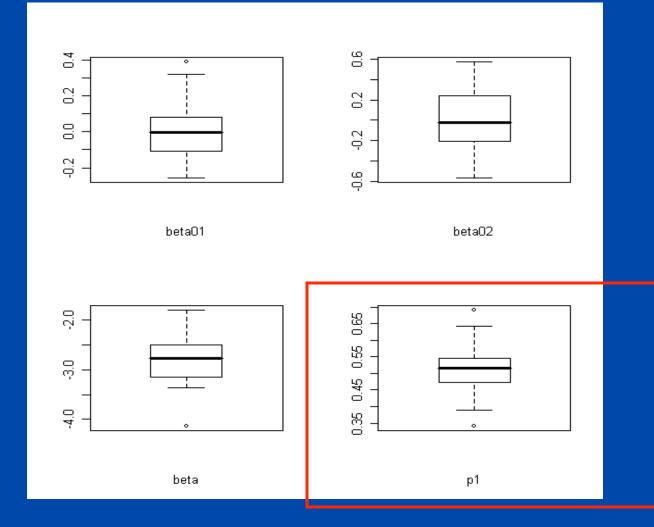
Simulation Study Diff. Meas.; $\beta_{12} = \beta_{13} = -1.4$

Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β ₁₂	-1.61	0.32	-2.00	0.31
β ₁₃	-0.08	0.28	-1.02	0.34

Simulation Study Non-diff meas; $\beta_{12} = \beta_{13} = -1.4$

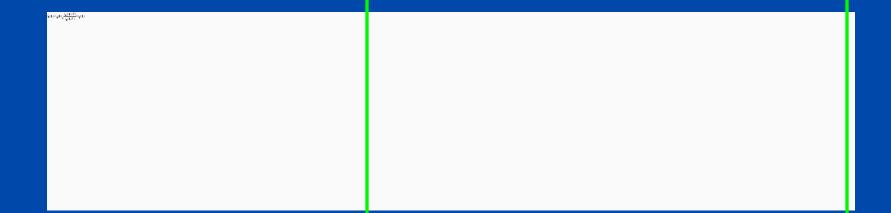
Param.	Penalized		LCR	
	Estimate	SE	Estimate	SE
β ₁₂	-1.45	0.34	-1.45	0.30
β ₁₃	-1.38	0.31	-1.38	0.31

Simulation Study Non-diff meas; $\beta_{12} = \beta_{13} = -1.4$



One empirical lead Deciding the extent of penalization

• Notice the form of $F_{V|x}^*(v|x_i)$:



• Idear1: y Right penalty yields $f^* = f$

Simulation study Three-class model

- Small: 100 reps; single x~Unif(-.5,.5)
- Multiple n: Here, =2000
- Poly Log Reg: $\beta_{01} = \beta_{02} = 0$; $\beta_{12} = -1.4$; $\beta_{12} = -2.8$
- Measurement:

Class 1	Class 2	Class 3
.15	.85	.85
.15	.85	.85
.15	.85	.85
.15	.13	.85
.15	.13	.85

Simulation study Three-class model

- Two scenarios (among more)
 Frank LCR
 - <u>Differential measurement</u>: last two items have increased log(odds =1) per unit x of 1.4 within each class
- Premise: $f_{V|x}^{*}(v|x_{i},\theta)$, $f_{V|x}(v|x_{i},\theta)$ quite different
- Measure: Kullback-Leibler distance

KL Distance: f*, f Scenario 1, n=2000

-3.4 -3.3 -3.2 -3.1 -3.0 -2.9 -2.8 -2.7 -2.6 -2.5 -2.4 -2.3 -2.2 True -2.0 4.99 4.76 4.76 4.86 4.89 5.15 5.26 5.42 6.23 6.34 6.93 7.59 7.99 -1.9 4.58 4.28 4.40 4.57 4.19 4.42 4.62 5.09/5.15 5.62 6.03 6.91 7.31 -1.8 4.52 4.36 4.18 4.07 3.88 3.96 4.22 4.26 4.55 5.09 5.52 5.96 6.58 -1.7 4.30 4.05 3.90 **3.64** 3.85 **3.71 3.73 4**.05 4.35 4.46 4.92 5.33 5.77 -1.6 4.56 4.21 3.80 3.62 3.52 3.54 3.67/3.69 3.88 4.07 4.36 4.88 5.46 -1.5 4.67 4.11 3.88 3.70 3.56 3.41 3.46 3.42 3.75 3.74 4.28 4.52 4.85 -1.4 4.87 4.39 3.91 3.84 3.62 3.27 3.62 3.40 3.69 3.68 3.70 4.03 4.52 -1.3 5.25 4.73 4.50 4.16 3.86 3.54 3.45 3.46 3.39 3.52 3.78 4.12 4.43 -1.2 5.58 4.99 4.76 4.47 4.16 3.81 3.70 3.60 3.75 3.74 3.85 4.25 4.30 -1.1 6.25 6.05 5.26 4.90 4.55 4.14 4.20 4.03 4.01 3.94 3.91 4.45 4.28

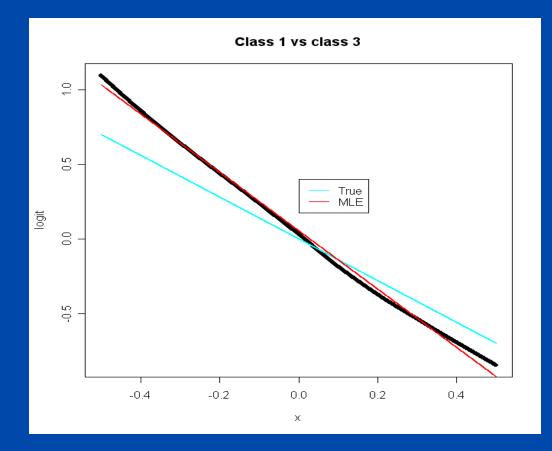
KL Distance: f*, f Scenario 2, n=2000

 $\hat{\beta}_{22}$ -3.8 -3.7 -3.6 -3.5 -3.4 -3.3 -3.2 -3.1 -3.0 -2.9 -2.8 -2.7 -2.6

 \hat{eta}_{12} -2.4 4.03 4.37/4.63 5.05 5.39 5.93 6.35 7.17 8.00 8.76 9.36 10.40 11.74 -2.3 3.79 3.87 4.10 4.59 4.93 5.14 5.84 6.38 6.76 7.79 8.55 9.46 10.50 -2.2 3.48 3.63 3.90 3.98 4.27 4.60 5.20 5.76 6.17 7.01 7.78 8.26 9.65 -2.1 3.31 3.17 3.47 3.51 3.95 4.25 4.69 5.04 5.64 6.34 7.01 8.09 9.07 -2.0 **3.19 3.29 3.41 3.33 3.70 3.94 4.34 4.60 5.10 5.62 6.70 7.24 8.02** -1.9 3.17 3.09 3.19 3.27 3.39 3.64 3.99 4.25 4.93 5.40 6.17 6.90 7.37 -1.8 **3.31 3.24 3.22 3.26 3.35 3.63 3.98** 4.35 4.75 5.12 5.34 6.40 7.00 -1.7 **3.56 3.33 3.43 3.32 3.31 3.57 3.85** 4.17 4.40 4.79 5.43 6.00 6.33 6.62 -1.6 3.83 3.77 **3.60 3.69 3.68 3.62** 3.80 4.19 4.65 4.87 5.38 6.21 -1.5 4.36 3.95 4.02 3.97 3.89 3.82 4.05 4.24 4.56 5.05 5.37 5.86 6.36 -1.4 4.90 4.69 4.43 4.28 4.34 4.46 4.35 4.65 4.88 5.11 5.41 5.99 6.49 -1.3 5.56 5.41 5.11 4.95 4.77 4.84 4.72 4.74 5.01 5.49 5.85 6.19 6.60 -1.2 6.41 5.97 5.87 5.59 5.37 5.17 5.33 5.18 5.52 5.96 6.08 6.31 6.99

True

Simulation Study Empirical support for "penalty"?



 Average conditional probability estimates amazingly stable

 Distinction: Y|V*,x

Frailty analysis: Data InCHIANTI (*Ferrucci et al., JAGS, 48:1618-25*)

- Aim : Causes of walking decline
- Brief design
 - Random sample \geq 65 years (n=1270)
 - Enrichment for oldest-old, younger ages
 - Participation: > 90% in the primary sample
 - Home interview, blood draw, physical exam
- Dysregulation: inflammation 7 cytokines
 - IL-6, CRP, TNF-a, IL-1RA, IL-18, IL-1B, TGF- β
 - Here: concern = poorer inhibition
- Frailty: Fried criteria (as before)

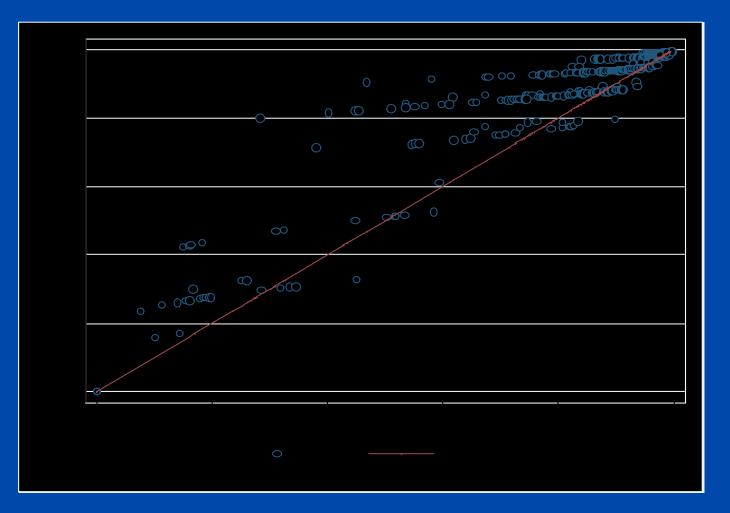
Frailty analysis: Results

- Measurement model: 2 classes
 - Conditional probabilities similar to WHAS
 - Lower "frail" prevalence (15% vs. 27%)
- Regression model
 - 1 SD worse inhibition index associated with 35% reduction in non-frail odds (z ~ 3)
 - Regression coefficient on original index scale: 3.00
- Next: Vary regression coefficients in increments of +/- 0.5, up to +/- 2.0

Frailty analysis: Results <u>Posterior probs. from different fits</u>



Frailty analysis: Results Posterior probs. non-frail, different fits



Frailty analysis: Results Age-adjusted relation to mobility

Frailty fit: inflam. slope	Mobility slope (frail vs non)	SE
ML – 2.0	-1.1	.089
ML - 1.0	-1.0	.087
ML – 0.5	-1.0	.086
ML	-0.99	.085
ML + 0.5	-0.93	.085
ML + 1.0	-0.92	.085
ML + 2.0	-0.82	.083



- Presented: Frameworks for measurement

 of complex geriatric health states
 incorporating biological knowledge
- Demonstrations
 - Frailty in WHAS
 - Frailty and inflammatory dysregulation in In CHIANTI

Rationale for the proposal

vs looser internal validation criteria?
 – estimability

vs Bayesian approach

 depends on degree of empiricism
 if balance by "consensus"—Bayesian

Allows some distrust of the data

Research needed

- Theory elicitation, incorporation
- Methodology freeing measurement model estimation to "move" with "penalty"
 - Rotation?
 - Penalty on conditional probabilities
- Compromise of latent variable, predictive approaches
- Best index derivation

Implications

- Refined understanding of aging states and their measurement
 - Integrating biology
 - Increasing sensitivity, specificity
- Heightened accuracy, precision for

 Delineating etiology
 Developing and targeting interventions

Acknowledgments

- Hopkins Colleagues
 Linda Fried, Ron Brookmeyer, Yi Huang,
 Jeannie-Marie Leoutsakos, Jeremy Walston,
 Qian-Li Xue, Scott Zeger
- Colleagues outside of Hopkins Luigi Ferrucci, Jack Guralnik, Don Ingram, Richard Miller
- Funding / Institutional Support Johns Hopkins Older Americans Independence Center, National Institute on Aging, Alliance for Aging Research